# **Systematic Review**

# High Prevalence of Neuropathic Pain Component in Patients with Low Back Pain: Evidence from Meta-Analysis

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**Background:** Low back pain (LBP) is a complex syndrome which includes a nociceptive (NcP) component, a neuropathic (NeP) component, or a mixture of components (mixed pain). The NeP component (NePC) in LBP is defined as the presence of NeP with or without an NcP.

**Objective:** This meta-analysis aimed at assessing the pooled prevalence of NePC in patients with LBP and at identifying the factors causing significant heterogeneity in reported prevalence.

Study Design: Meta-analysis.

**Methods:** A systematic literature search was carried out, with inclusion of all epidemiological studies describing the NeP prevalence levels in LBP patients while using standard diagnostic methods. The "pooled prevalence rate (PPR)" of NePC, either on its own or in combination with NcP, was calculated. A pre-specified subgroup analysis was carried out, considering LBP duration, presence of leg pain, diagnostic method(s), and questionnaire(s) used.

**Results:** The meta-analysis included 20 studies relating to a total of 14,269 LBP patients, of whom 7,969 patients (55.8%) were identified as presenting with NePC. The pooled PR (95% CI) of NePC in patients with LBP was 0.47 (0.40 - 0.54), while the pooled PR of NcP was 0.56 (0.48 - 0.63). Higher NePC pooled PR values were identified in LBP with leg pain as compared to uncomplicated LBP (respectively: 0.60; 0.47 - 0.73 vs 0.27; 0.23 - 0.31;  $P_{\text{interaction}}$  < 0.01).

**Limitations:** The quality of the included studies was assessed using ad-hoc criteria. Due to the limited number of available studies, one may need to be cautious in reaching conclusions about the impact of disease duration on NePC prevalence values. We pooled studies which used a range of different diagnostic methods, with putatively different sensitivity/specificity diagnosing levels.

**Conclusion:** Overall, high NePC prevalence levels were here identified in LBP patients. As the pain is a subjective phenomenon and there is no gold standard for the diagnosis of NePC, there is the possibility that the pooled effect estimate may alter depending upon the diagnostic method used.

**Key words:** Neuropathic pain, nociceptive pain, low back pain, symptom-based questionnaire, chronicity

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ow back pain (LBP) is the most frequently observed chronic pain condition (1). It may be associated with leg pain and/or motor/ sensory/reflex deficits in the nerve root distribution areas. Thus, LBP can present with nociceptive (NcP),

neuropathic (NeP), or mixed pain components (mixed pain) (1). The neuropathic pain component (NePC) in LBP is defined as the presence of NeP either on its own or associated with NcP. Recently, the development of multiple measurement scales, such as the

painDETECT questionnaire (PDQ) (2), the neuropathic pain questionnaire (NPQ) (3), the Leeds Assessment of Neuropathic Symptoms and Signs (LANSS) (4), the ID-pain (5), the Douleur neuropathique 4 (DN4) (6), and the standardized evaluation of pain (StEP) (7) have facilitated a better identification of NePC in LBP conditions. The extent of involvement of NePC in LBP patients is extensively debated, with reported prevalence rates being in the range of 28.1% to 71.2% (8,9). Recognition of NePC is important, since it may significantly affect the patient's psychological and social functioning, while placing considerable financial demands on society (10).

Categorizing LBP as either NcP or NeP is not always straightforward (11). A range of different approaches have been proposed to facilitate this distinction, including physician's judgement, International Association for the Study of Pain (IASP) Treed's criteria (12), different questionnaires (2-7), and finally quantitative sensory (QST) and pharmacological testing methods (13). Identifying the presence of NePC in LBP is important, since this may present as a challenging treatment/management issue (14), typically associated with low rates of treatment success (15). Patients with acute LBP may show NeP symptoms similar to those of chronic LBP (CLBP), with lack of recognition possibly leading to suboptimal drug dosages being administered, with the risk that the acute pain condition may persist and become subacute/ chronic (16,17).

The current meta-analysis is aimed at assessing the pooled prevalence of NePC in patients with acute, sub-acute, and chronic LBP and identifying the factors causing significant heterogeneity in reported prevalence.

## **METHODS**

# **Search Strategy**

Two electronic databases (PubMed and EMBASE, from their inception until August 2015) were searched independently by 2 authors without any language restriction. Both full texts and MeSH terms were considered to identify those epidemiological (e.g., case control and cohort) studies focusing on assessing NeP prevalence levels in LBP patients. Searches were conducted using synonyms and combinations of the following search terms: "neuropathic pain," "prevalence," "low back pain," and "leg pain." In addition, related systematic reviews and selected studies' reference lists were screened for any other relevant study to be included in the present analysis.

## Study Selection Criteria

Two authors independently reviewed the titles and abstracts of all identified citations as per inclusion and exclusion criteria. An abstract was considered to be relevant if it reported original data, originated from epidemiological studies, and reported the prevalence of NePC in patients with LBP. For the present analysis, patients were classified in 3 groups: a) patients with exclusive NcP, b) patients with exclusive NeP, and c) patients with mixed pain (e.g., suffering from both NeP and NcP). NePC was here defined as presence of NeP symptoms exclusively and/or associated with NcP symptoms.

Studies were excluded if data regarding prevalence of NeP in LBP patients were missing/not clearly reported, and when diagnostic methods were either not standardized or not reported. Single case reports, comments, letters, editorials, protocols, guidelines, and review papers were excluded. When there were multiple publications for the same population, only the most detailed report was included. Any discrepancies were resolved by a consensus conference.

# **Data Extraction and Synthesis**

Two authors independently extracted a range of relevant information, including first author name, year of publication, country of patients' recruitment, study setting, study design, sample size, diagnostic methods used to assess NePC, NeP/NePC reported prevalence rates (PR) and 95% confidence interval (CI), duration of disease, and definitions of NeP, NcP, and NePC. If unavailable, both PR and 95% CI of NeP/NePC were calculated from the raw data available. If a study reported PRs of NeP and mixed pain separately, they were pooled together to get an estimate of NePC.

## **Quality Assessment**

Consistent with the approach of both Hoogen-doorn et al (18) and De Vet et al (19), all included studies were independently assessed for their methodological quality by 2 authors. These researchers had developed and tested a list of 23 criteria to be used to assess the methodological quality of prospective, cohort, and case control studies. Out of this list of criteria, Fishbain et al (20) selected 10 criteria that were deemed appropriate for the assessment of NeP prevalence in LBP. To this list, Fishbain et al (20) have added one extra criterion. We have excluded 5 questions and added one (relating to the characteristics of recruitment, e.g., consecutive or not) question for this study, resulting in a total of 7 criteria, including study design related quality assessment

measures like clearly defined objective, standard data collection procedure, and statistical analysis plan. For each of the included studies, each criterion was rated either fulfilled (+), unfulfilled (-), or not applicable (NA). A quality score for each paper was obtained by counting the number of positives obtained; the score was then divided by 7 (minus the number of NAs) and multiplied by 100, providing the percentage quality score. Studies scoring less than 60% were excluded for further review. Any discrepancy in quality assessment was discussed and internally resolved between 2 authors.

# **Data Synthesis and Analysis**

As the primary outcome measure, we calculated the weighed pooled PR of NePC in patients with LBP. The weight, Wi, assigned here to each study was the inverse of the estimated variance. Where different prevalence estimates in the same patient group, using different diagnostic methods, were reported those prevalence estimates resulting from the most reliable method were chosen. For the present meta-analysis, the pooled PR, with 95% CI, was reported as a summary measurement, with the results being presented as a forest plot. Cochran Q and I<sup>2</sup> statistics were used to assess heterogeneity, with a Q statistic P value of < 0.10, and an  $I^2$  value >50% having been considered as statistically significant. If heterogeneity was identified, the DerSimonian-Laird random-effects model was applied; otherwise, the fixed effects-model (inverse variance) was considered (21). A pre-specified subgroup analysis was carried out to assess the source of heterogeneity, according to: i) diagnostic method(s) for identification of NeP (e.g., physician- or questionnaire-based), ii) duration of LBP (e.g., acute, subacute, or chronic), iii) presence of associated leg pain, iv) primary objective of the study being/not being the assessment of NeP prevalence, v) number of study centers involved (e.g., single vs multi-center), and vi) questionnaires used for the diagnosis (e.g., LANSS, DN4, PDQ, and NPQ). Tests for interaction using summary estimates were performed, using the method described by Altman and Bland (21).

To evaluate the stability of our results, we also performed a one-way sensitivity analysis. The scope of this analysis was to evaluate the influence of individual studies by estimating the average PR in the absence of each study. The present work was performed in accordance with the guidelines proposed by the Metanalysis of Observational Studies in Epidemiology group and Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA; supplementary material 1)

(22). The publication bias was assessed using both the funnel plot and the Begg's test. The Duval and Tweedie nonparametric trim-and-fill method was performed to further assess the effect of potential publication bias, a method which negotiates the lack of studies on a particular side of the funnel (23). When data were not uniformly reported to allow a formal statistical analysis, we presented the available data in a narrative format. All statistical tests were 2-sided, and a P value < 0.05 was considered statistically significant, except where otherwise specified. Data were analyzed using STATA version 11.0 (StataCorp, College Station, TX, USA).

# RESULTS

#### **Search Results**

After applying the initial search strategy, a total of 160 citations were identified (Fig. 1). After reviewing the citations, 124 of them were considered ineligible, as they were reviews/editorials/case reports, hence not meeting the inclusion criteria. In examining the full text of the remaining 36 studies and reviewing the related reference lists, 5 further studies were considered. After a detailed evaluation of these 41 manuscripts, 21 of them were excluded for a range of reasons, which are better detailed in Fig. 1. As a result, 20 studies were included in the present meta-analysis.

# **Description of Included Studies**

All the studies had been published between 2004 and 2015. Most (n = 10; 50%) originated from Europe (2,7,9,24-30), followed by Asia (n = 5; 25%) (10,17,31-33), and Africa (n = 4; 20%) (8,34-36). The study sample recruitment site was not reported in one study (37).

Out of these studies, 17 (85%) were cross sectional, 2 were cohort, and one was a case control study. Eleven (55%) studies were multi-center (2,10,24-27,29-31,34,35) whereas the remaining were single-center studies (7-9,17,28,32,33,36,37). Twelve (60%) studies were carried out primarily to assess the prevalence of NeP in patients with LBP, while the remaining focused either on the effects of NeP on a range of patient-reported outcomes or on the validation of a range of NeP-related questionnaires. Eighteen (90%) studies included patients with LBP as the main clinical concern, while 2 studies included patients with multiple pain conditions. The LBP etiology was reported in 8 (40%) studies (10,17,24,31,32,34-36).

All studies were carried out in hospital settings. Most (n = 16) studies had reported a diagnosis of

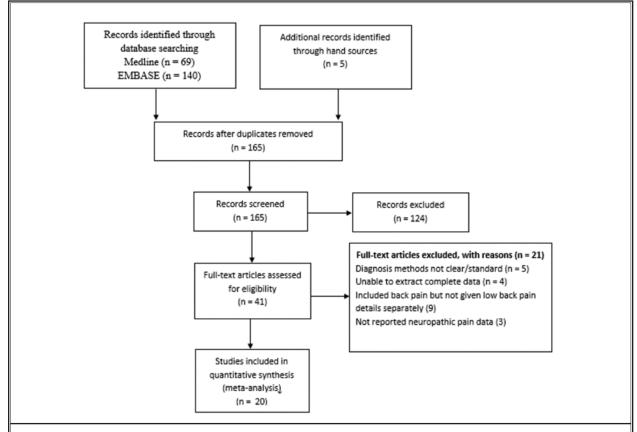


Fig. 1. Flow chart showing the number of citations retrieved by individual searches and the number of studies included in the review.

NePC which relied solely on symptom-based questionnaires, while the sole physician's criteria were used in 2 studies (5,27). Only one study (7) reported an NePC diagnosis made with the help of both physician's criteria and symptom-based questionnaires. Most commonly used questionnaires included PDQ (n = 7 studies) (2,9,17,25,26,28,33), followed by DN4 (6,8,24,36,37), and LANSS (6,31,32,34,35) (n = 5 each). Conversely, NPQ (n = 2 studies) (10,33) and StEP (n = 1 study) (7) were less frequently considered (Table 1).

## **Quality Assessment**

All the studies had a quality score ranging from 71% to 100%, hence were eligible for inclusion. The average (SD) score was 92.8% (9.8%) (Table 2). Two studies presented with a score of 71.4%; in fact, they were not characterized by an adequate sample size (e.g., n < 50 patients) and had not considered a consecutive patient recruitment (10,33).

### **Primary Outcome**

As a significant heterogeneity was found between studies (Cochrane Q  $\leq$  0.01, I<sup>2</sup> = 98), a random effects model was chosen (Table 3). The present meta-analysis included 14,269 patients with LBP, of whom 7,969 patients (55.8%) were identified with NePC. The pooled PR (95% CI) of NePC in patients with LBP out of all the included studies was 0.47 (0.40 - 0.54) (Fig. 2), with pooled PR for NePC being in the 0.28 to 0.71 range.

## **Sensitivity Analysis**

The estimated effect size did not deviate much by excluding any study, one at a time, with pooled PR values ranging between 0.42 and 0.49.

### **Publication Bias**

The visual examination of the funnel plot revealed only minimal levels of asymmetry (Fig. 3). This was further confirmed by the Begg's test (P = 0.58) which indicated little or no publication bias. The NePC pooled

Table 1. Characteristics of 20 studies included in the analysis to assess the prevalence rate of the neuropathic pain component in patients with low back pain.

Author, year (Reference number of the study)	Country	Method of Diagnosis^	Questionnaire	Sample Size	Duration of LBP#	Prevalence (95% CI)	Study Quality <sup>\$</sup> %	
Hassan et al, 2004 (34)	SA	1	LANSS	100	Chronic	41 (31.4 – 50.6)	83.3	
Kaki et al, 2005 (35)	SA	1	LANSS	1,169	Chronic	54.7 (51.8 – 57.5)	100	
Freynhagen et al, 2006 (2)	Germany	1	PDQ	7,772	Chronic	64.7 (63.6 – 65.8)	100	
Freynhagen et al, 2006 (27)	Germany	2	-	717	Chronic	33 (29.6 – 36.4)	100	
Scholz et al, 2009 (7)	UK	1,2*	StEP	137	Chronic	54.7 (46.4 - 63)	85.71	
Sissi et al, 2010 (31)	UAE	1	LANSS	1,134	Chronic	55.4 (52.5 – 58.3)	100	
Attal et al, 2011 (24)	France	1	DN4	132	Chronic	70.4 (62.6 – 78.2)	100	
Beith et al, 2011 (25)	England, UK	1	PDQ	343	NR	40.5 (35.3 – 45.7)	85.71	
Morsø et al, 2011 (28)	Denmark	1	PDQ	145	Chronic	45.5 (37.4 – 53.6)	100	
Ouedraogo et al, 2011 (36)	Burkina Faso, Africa	1	DN4	107	Acute/Subacute, Chronic	49.5 (40 – 59)	100	
Smart et al, 2012 (29)	Ireland, UK	2	-	464	Acute/Subacute, Chronic	44.8 (40.3 – 49.3)	83.33	
				123	Acute/Subacute	32.5 (24.2 – 40.8)		
				341	Chronic	49.3 (44 – 54.6)		
Walsh et al, 2012 (30)	Ireland	1	S-LANSS, DN4*	45	NR	42 (27.6 – 56.4)	85.71	
Uher et al, 2012 (9)	Czech Republic	1		66	Acute/Subacute, Chronic	71.2 (60.3 – 82.1)	100	
			PDQ	40	Acute/Subacute	70 (55.8 – 84.2)		
				26	Chronic	73 (56 – 90.1)		
Yamashita et al, 2013 (10)	Japan	1	NPQ	17	Chronic	29.4 (7.7 – 51)	66.66	
Forster et al, 2013 (26)	Germany	1	PDQ	1,083	NR	30.7 (28 – 33.4)	83.33	
Selimoglu et al, 2013 (37)	NR	1	DN4	224	Chronic	55.8 (49.3 – 62.3)		
Doualla et al, 2013 (8)	Cameroon	1	DN4	167	Chronic	28.1 (21.3 – 35)		
Hiyama et al, 2014 (17)	Japan	1		331	Acute/Subacute, Chronic	42.5 (37.2 – 47.8)	100	
			PDQ	201	Acute/Subacute	25.3 (19.3 – 31.3)		
				130	Chronic	31.5 (23.5 – 39.5)		
Sakai et al, 2015 (33)	Japan	1	NPQ*, PDQ	30	Chronic	43.3 (25.6 – 61)	66.66	
Park et al, 2015 (32)	Korea	1	LANSS	86	NR	36 (25.8 – 46.1)	100	

NR, Not reported; SA, South Africa; UK, United Kingdom; UAE, United Arab Emirates; LANSS, Leeds Assessment of Neuropathic Symptoms and Signs; NPQ, Neuropathic Pain Questionnaire; PDQ, PainDETECT Questionnaire; DN4, Douleur Neuropathique 4 questionnaire

PR value (e.g., 0.47; 0.40 - 0.54) obtained using the random effects model remained unchanged using the Duval and Tweedie's (23) trim-and-fill analysis.

## **Sub-group Analysis**

With the physician-diagnosis approach, the pooled

PR of NePC in LBP patients was lower (0.44; 0.32-0.55) than the pooled PR resulting from the questionnaire-based diagnosis (0.48; 0.40-0.55;  $P_{\rm interaction}=0.58$ ). Similarly, the pooled PR of NePC was only marginally higher (0.49; 0.41-0.57) in CLBP patients in comparison with patients affected by acute/subacute LBP (0.45; 0.28-0.58)

<sup>^ 1:</sup> questionnaire; 2: clinician rated \* Standard method considered if more than one method used for the diagnosis of NPC

<sup>#</sup> Acute: duration of LBP less than 6 weeks; Subacute: duration of LBP 6 to 12 weeks; Chronic: duration of LBP more than 12 weeks. C.I; Confidence Interval

<sup>§</sup> Quality was assessed using scale developed by Fishbain et al (here modified)

8 Hiyama et al, 9 Hassan et al, 10 Ouedraogo et al, 11 Sissi et al, 12 Scholz et al, 13 Walsh

20 Doualla et al, NA Not Available.

al, 16 Forster et al, 17 Sakai et al, 18 Park et al, 19 Selimoglu et al,

100 + + 20 85.71 + + 19 001 28 + + + 71.40 + 85.71 + + + + 16 001 2 + + + + + + 100 + + + + + + 7 85.71 + + + 133 85.71 + 2 100 + + +  $\equiv$ + + + 100 2 + + + + + 85.71 + + + + 6 001 + + + + + + 00 100 + + + + + + 85.71 + + + 9 71.40 + + + LO 100 + + + 100 + + + + + + ಣ 85.71 a 001 + + + + standardized methods of acceptable quality Positive if data were collected by means of Smallest group larger than 50 participants statistical analysis was appropriate for the Positive if the recruitment of participants Positive if the study had a clearly defined the population according to age and sex) Positive if the main features (description of the study population were adequately of the sampling frame; distribution of Positive if the methods used for the Restricted to a homogenous study for neuropathic pain was consecutive population described Criteria Score in %\* S.No. 5 9

\* Score ranges from 0% to 100%, where higher percentages represent higher quality of the study +, Study fulfilled the criterion; -, study unfulfilled the criterion; NA, not applicable.

I Attal et al, 2 Beith et al, 3 Kaki et al, 4 Morsø et al, 5 Yamashita et al, 6 Smart et al, 7 Freynhagen et al,

al, 14 Freynhagen (2) et al, 15 Uher et

0.62;  $P_{\rm interaction} = 0.71$ ). Conversely, significantly higher pooled PR values of NePC were associated with leg pain (0.60; 0.47 – 0.73) as compared to those conditions without an associated leg pain (0.27; 0.23 – 0.31;  $P_{\rm interaction} < 0.01$ ). When the subgroup analysis was performed according to the type of screening questionnaires used, comparable PR were obtained with LANSS (0.47; 0.29 – 0.64), DN4 (0.49; 0.33 – 0.65), and PDQ (0.48; 0.34 – 0.63). However, relatively lower PR values of NePC were identified when the NPQ was used (0.36; 0.90 – 0.64).

# **Secondary Outcomes**

The pooled PR of mixed pain in patients with LBP (n = 7 studies) was found to be 0.28 (0.23 – 0.34) (Table 4), while the pooled PR of sole NeP in patients with LBP (n = 7 studies) was found to be 0.23 (0.11 – 0.34). The pooled PR of NcP in patients with LBP in all studies was 0.56 (0.48 – 0.63).

### DISCUSSION

The present meta-analysis of 20 studies, including 14,269 participants, found that the PPR of NePC in patients with LBP was 47% (40% - 54%), hence NePC may need to be considered as a significant clinical problem in these patients. One of the major strengths of the current study is that we focused on both the neuropathic and the nociceptive components, while considering not only the chronic, but also the acute/ subacute, LBP conditions. Moreover, previous studies considered mathematic calculations only to estimate NePC pooled PR values, as opposed to the formal meta-analysis approach which was carried out here. The NePC figure (47%) identified is consistent with previous findings, e.g., Morso et al (45.5) and Ouedraogo et al (49.5) (28,36), but it is at odds with the figures reported by both Fishbain et al (36.6%) and Mehra et al (90.4%) (20,38). The studies here reported a prevalence of NePC ranging from 28.1% to 71.2%, with these levels

Table 2. Quality assessment of the studies included

Table 3. Pooled prevalence rates of NePC in different subgroups of patients with LBP.

		Neuropathic pain component				Heterogeneity					
Study Characteristics		Studies	No. of Patients	PR (95% CI)	P value	Cochrane Q	$I^2$	P Subgroup#			
All studies	14269	20	7969	0.47 (0.40 - 0.54)	< 0.01	< 0.01	98.0	NA			
Leg pain											
Present	572	5	300	0.60 (0.47 - 0.73)	< 0.01	< 0.01	89.9	< 0.01			
Absent	479	2	89	0.27 (0.23 – 0.31)	< 0.01	< 0.01	97.5				
Diagnostic method											
Physician-based	1318	3	520	0.44 (0.32 – 0.55)	< 0.01	< 0.01	93.6	0.58			
Questionnaire-based	12951	17	4792	0.48 (0.40 - 0.55)	< 0.01	< 0.01	97.9				
Duration of pain											
Acute/Subacute	364	3	80	0.45 (0.28 - 0.62)	< 0.01	< 0.01	90.5	0.71			
Chronic	51666	15	40585	0.49 (0.41 - 0.57)	< 0.01	< 0.01	97.6				
Study Centres	Study Centres										
Single center	1293	09	598	0.47 (0.39 – 0.55)	< 0.01	< 0.01	87.8	0.53			
Multi-center	12976	11	7371	0.46 (0.37 - 0.56)	< 0.01	< 0.01	98.8				
Primary objective was to assess the prevalence of NePC in LBP											
Yes	5267	12	2373	0.44 (0.36 - 0.52)	< 0.01	< 0.01	96.6	0.57			
No	9002	08	5596	0.51 (0.41 - 0.60)	< 0.01	< 0.01	96.1				
Questionnaire used											
LANSS	2489	4	1339	0.47 (0.29 - 0.64)	< 0.01	< 0.01	85.1				
DN4	675	5	337	0.49 (0.33 – 0.65)	< 0.01	< 0.01	94.2				
PDQ	9740	6	5755	0.48 (0.34 - 0.63)	< 0.01	< 0.01	99.1				
NPQ	47	2	18	0.36 (0.90 – 0.64)	< 0.01	< 0.01	0.00				

Comparison between subgroups

NA, Not Available; CI, Confidence Interval; PR, Prevalence Rate; LANSS, Leeds Assessment of Neuropathic Symptoms and Signs; NPQ, Neuropathic Pain Questionnaire; NePC, Neuropathic pain component; PDQ, PainDETECT Questionnaire; DN4, Douleur Neuropathique 4 questionnaire; LBP, low back pain.

of variations possibly being explained by differences in inclusion criteria, sample size, study settings, and diagnostic methods used (8,9). Furthermore, some diagnostic questionnaires fail to identify the mixed pain conditions, hence underestimating the presence of NeP symptoms (31,35,36). According to the current sensitivity analysis, excluding any single study did not have here any major impact on the combined results, which may well suggest that the present meta-analysis results are pretty consistent. Furthermore, both a robustness of results and lack of publication bias were confirmed.

Overall, a physician/clinician decision is considered to be the gold standard in diagnosing any type of pain, with the diagnostic accuracy of screening questionnaires not always being considered satisfactory (39). However, similar NePC prevalence levels were identified in those studies using questionnaires (0.48; 0.40 – 0.55)

vs those considering the clinician-based diagnosis (0.44; 0.32 - 0.55).

LBP duration was hypothesized to be a possible factor associated with different NePC prevalence levels. However, the NePC PPR value was only marginally higher in patients with CLBP than in those experiencing acute/subacute LBP (respectively: 0.49; 0.41 – 0.57; vs 0.45; 0.28 – 0.62; ns). Indeed, acute and chronic LBP cases differ in their etiology and pathophysiology, with acute LBP being usually a self-limiting, protective, biological function of on-going tissue damage. Acute pain is mostly nociceptive in origin and occurs secondary to stimulation of A-delta and C-polymodal pain receptors (16). Conversely, CLBP is in itself a disease process which can persist for years after the initial injury and which can be refractory to multiple treatment modalities. Inadequately treated acute pain conditions may be asso-

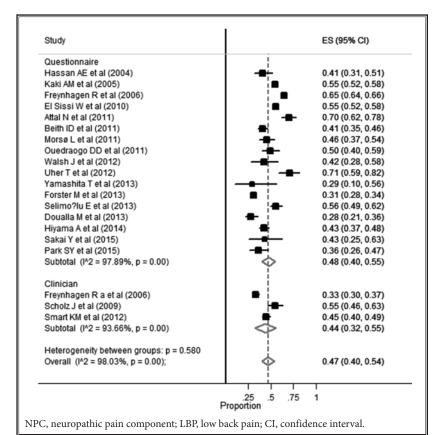


Fig. 2. Pooled estimate of prevalence rate and 95% CIs of NPC in patients with LBP in all the included studies (20 studies) comprising 14,269 patients with LBP. The pooled prevalence rate of NPC in patients with LBP in all the included

studies (20 studies) was found to be 0.47 (at 95% CI, 0.40 - 0.54).

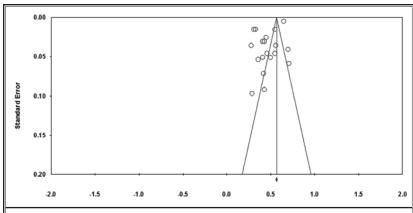


Fig. 3. Funnel plot to assess the publication bias (effect estimates on the horizontal scale and study size on the vertical axis).

ciated with NeP symptoms (16) and may become chronic. Consistent with this, a higher prevalence of NePC is typically identified in CLBP as opposed to acute/subacute LBP (17).

The presence of leg pain was hypothesized as another possible factor associated with different NePC prevalence levels. As expected, and consistent with previous findings (17), the NePC prevalence was significantly higher in patients affected by leg pain than in those with uncomplicated LBP. Although both NeP and NcP mechanisms are thought to contribute to leg pain in LBP, neuropathic mechanisms play a more significant role (1,27) Evidence also suggests that patients experiencing a radiating leg pain may complain of intense pain and significant disability (40), delayed recovery, and a tendency to chronicity (41). Furthermore, they frequently require surgical interventions for their LBP to be managed (42). To better identify the presence of NePC in LBP, some questionnaires (e.g., PDQ) have been designed to specifically include a question about the presence of leg pain (12). Results of the present meta-analysis might not be generalizable to patients with failed back syndrome where the prevalence of NeP is more common and the cause for NeP is quite different.

### Limitations

The quality of the studies selected was assessed using an ad hoc, customized scale. Although not specifically validated, the scale presented has been designed while including further assessment criteria to previous instruments. Due to the limited number of available

studies, one may need to be cautious in concluding about the impact of disease duration on NePC prevalence values. Furthermore, the quality of data collected from those studies whose primary aim was not the assessment of NeP prevalence may be questionable. Similarly, one could wonder about the appropriateness of grouping together studies which used a range of different diagnostic methods, with putatively different sensitivity/specificity diagnosing levels. To address this issue, however, a subgroup analysis based on type of questionnaires used was carried out. The limitation of current diagnostic methods is that we cannot quantify how much neuropathic there is vs. nociceptive to direct the clinician to the appropriate therapy. The studies included did not present with sufficient data to perform a subgroup analysis to estimate NeP prevalence levels

according to specific clinical conditions/pathologies involved. Finally, only a few studies included patients with LBP with an associated leg pain condition (9,28,30), so that the results presented may need to be interpreted with some caution. There are limited studies from the US.

Given the difficulties in treating and managing clients affected by low back pain with a neuropathic component further research studies with larger sample sizes and increased diagnostic accuracy levels should be promoted. As pain is subjective phenomenon and as there is a lack of a gold standard for the diagnosis of NP, there is the possibility that the pooled effect estimate may alter if the diagnostic ability of the available and used methods warrant the gold accuracy.

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